Week 03 notes

Repeated trials

01 Theory

Repeated trials

When a single experiment type is repeated many times, and we assume each instance is *independent* of the others, we say it is a sequence of **repeated trials** or **independent trials**.

The probability of any sequence of outcomes is derived using independence together with the probabilities of outcomes of each trial.

A simple type of trial, called a **Bernoulli trial**, has two possible outcomes, 1 and 0, or success and failure, or T and F. A sequence of repeated Bernoulli trials is called a **Bernoulli process**.

- Write sequences like *TFFTTF* for the outcomes of repeated trials of this type.
- Independence implies

$$P[TFFTTF] = P[T] \cdot P[F] \cdot P[F] \cdot P[T] \cdot P[T] \cdot P[F]$$

• Write p = P[T] and q = P[F], and because these are all outcomes (exclusive and exhaustive), we have q = 1 - p. Then:

 $P[TFFTTF] \gg pqqppq \gg p^3q^3$

• This gives a formula for the probability of any sequence of these trials.

A more complex trial may have three outcomes, A, B, and C.

- Write sequences like *ABBACABCA* for the outcomes.
- Label p = P[A] and q = P[B] and r = P[C]. We must have p + q + r = 1.
- Independence implies

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P[ABBACABCA] \gg pqqprpqrp \gg p^4q^3r^2
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• This gives a formula for the probability of any sequence of these trials.

Let *S* stand for the *sum of successes* in some Bernoulli process. So, for example, "S = 3" stands for the event that the number of successes is exactly 3. The probabilities of *S* events follow a **binomial distribution**.

Suppose a coin is biased with P[H] = 20%, and *H* is 'success'. Flip the coin 20 times. Then:

$$P[S=3] \gg \gg {\binom{20}{3}} \cdot (0.2)^3 \cdot (0.8)^{17}$$

Each outcome with exactly 3 heads and 17 tails has probability $(0.2)^3 \cdot (0.8)^{17}$. The *number* of such outcomes is the number of ways to choose 3 of the flips to be heads out of the 20 total flips.

The probability of at least 18 heads would then be:

$$P[S \ge 18] \gg P[S = 18] + P[S = 19] + P[S = 20]$$

 $\gg \gg {\binom{20}{18}} \cdot (0.2)^{18} \cdot (0.8)^2 + {\binom{20}{19}} \cdot (0.2)^{19} \cdot (0.8)^1 + {\binom{20}{20}} \cdot (0.2)^{20} \cdot (0.8)^6$

With three possible outcomes, A, B, and C, we can write sum variables like S_A which counts the number of A outcomes, and S_B and S_C similarly. The probabilities of events like " $(S_A, S_B, S_C) = (2, 3, 5)$ " follow a **multinomial distribution**.

02 Illustration

≡ Example - Multinomial: Soft drinks preferred

Folks coming to a party prefer Coke (55%), Pepsi (25%), or Dew (20%). If 20 people order drinks in sequence, what is the probability that exactly 12 have Coke and 5 have Pepsi and 3 have Dew?

Solution

The multinomial coefficient $\begin{pmatrix} 20\\ 12, 5, 3 \end{pmatrix}$ gives the number of ways to assign 20 people into bins according to preferences matching the given numbers, C = 12 and P = 5 and D = 3.

Each such assignment is one sequence of outcomes. All such sequences have probability $(0.55)^{12} \cdot (0.25)^5 \cdot (0.2)^3$.

The answer is therefore:

$$\binom{20}{12,5,3} \cdot (0.55)^{12} \cdot (0.25)^5 \cdot (0.2)^3 \quad \gg \gg \quad \frac{20!}{12! \, 5! \, 3!} \cdot (0.55)^{12} \cdot (0.25)^5 \cdot (0.2)^3$$

Reliability

03 Theory

Consider some process schematically with **components in series** and **components in parallel**:



- Each component has a probability of success or failure.
- Event *W_i* indicates 'success' of that component (same name).
- Then $P[W_i]$ is the probability of W_i succeeding.

Success for a *series* of components requires success of *each* member.

- Series components *rely on each other*.
- Success of the whole is success of part 1 AND success of part 2 AND part 3, etc.

Failure for *parallel* components requires failure of *each* member.

- Parallel components represent *redundancy*.
- Success of the whole is success of part 1 OR success of part 2 OR part 3, etc.

For series components:

$$P[W] = P[W_1W_2W_3] = P[W_1] \cdot P[W_2] \cdot P[W_3]$$

For parallel components:

 $P[W^c] =$ "failure" $\gg \gg P[W_1^c W_2^c W_3^c]$

$$\gg \gg (1 - P[W_1])(1 - P[W_2])(1 - P[W_3])$$

If $P[W_i] = p$ for all components *i*, then:

- Series components: $P[W] = p^3$
- Parallel components: $P[W] = 1 (1 p)^3$

To analyze a complex diagram of series and parallel components, bundle each:

- pure series set as a single compound component with its own success probability (the product)
- pure parallel set as a single compound component with its own success probability (using the failure formula)

This is like the analysis of resistors and inductors.

04 Illustration

\equiv Example - Series, parallel, series

Suppose a process has internal components arranged like this:



Write W_i for the event that component *i* succeeds, and W_i^c for the event that it fails.

The success probabilities for each component are given in the chart:

1	2	3	4	5	
92%	89%	95%	86%	91%	

Find the probability that the entire system succeeds.

Solution

Discrete random variables

05 Theory

🕆 Random variable

A **random variable (RV)** *X* on a probability space (S, \mathcal{F}, P) is a function $X : S \to \mathbb{R}$.

So *X* assigns to each *outcome* a *number*.

• (!) The word 'variable' indicates that the RV outputs *numbers*.

Random variables can be formed from other random variables using mathematical operations on the output numbers.

Given random variables *X* and *Y*, we can form these new ones:

$$rac{1}{2}(X+Y), \qquad X\cdot Y, \qquad \cos X, \qquad X^2, \qquad ext{etc}$$

Suppose $s \in S$ is some particular outcome. Then, for example, (X + Y)(s) is by definition X(s) + Y(s).

Random variables determine events.

- Given $a \in \mathbb{R}$, the event "X = a" is equal to the set $X^{-1}(a)$
- That is: the set of outcomes mapped to *a* by *X*
- That is: the event "*X* took the value *a*"

Such events have probabilities. We write them like this:

 $P[X=a] \quad \gg \gg \quad P[X^{-1}(a)]$

This generalized to events where *X* lies in some range or set, for example:

$$P[a \leq X < b], \qquad Pig[X \in \{2,4,5,6,9\}ig]$$

The axioms of probability translate into rules for these events.

For example, additivity leads to:

$$P[X < 0] + P[X = 0] + P[0 < X \le 3] + P[3 < X] = 1$$

A discrete random variable has probability concentrated at a discrete set of real numbers.

- A 'discrete set' means finite or countably infinite.
- The distribution of probability is recorded using a **probability mass function (PMF)** that assigns probabilities to each of the discrete real numbers.
- Numbers with nonzero probability are called **possible values**.

B PMF

The PMF function for X (a discrete RV) is defined by:

$$P_X(k) := P[X = k]$$

for $k \in \mathbb{R}$ a possible value.

A continuous random variable has probability spread out over the space of real numbers.

• The distribution of probability is recorded using a **probability density function (PDF)** which is *integrated over intervals* to determine probabilities.

B PDF

The PDF function for *Y* (a CRV) is written $f_Y(x)$ for $x \in \mathbb{R}$, and probabilities are calculated like this:

$$Pig[a \leq Y \leq big] = \int_a^b f_Y(x)\,dx$$



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For any RV, whether discrete or continuous, the distribution of probability is encoded by a function:

B CDF

The **cumulative distribution function (CDF)** for a random variable *X* is defined for all $x \in \mathbb{R}$ by:

$$F_X(x) = P[X \le x]$$

Notes:

- Sometimes the relation to X is omitted and one sees just "F(x)."
- Sometimes the CDF is called, simply, "the distribution function" because:
- 🕛 The CDF works equally well for discrete and continuous RVs.
 - Not true for PMF (discrete only) or PDF (continuous only).
 - There are *mixed* cases (partly discrete, partly continuous) for which the CDF is *essential*.

The CDF of a discrete RV is always a stepwise increasing function. At each step up, the jump size matches the PMF value there.

From this graph of $F_X(x)$:



we can infer the PMF values based on the jump sizes:

$P_X(-1)$	$P_X(0)$	$P_X(1)$	$P_X(2)$	$P_X(3)$	$P_X(4)$
0	1/8	3/8	3/8	1/8	0

For a discrete RV, the CDF and the PMF can be calculated from each other using formulas.

PMF from CDF from PMF

Given a PMF $P_X(x)$, the CDF is determined by:

$$F_X(x) = \sum_{k_i \leq x} P_X(k_i)$$

where $\{k_1, k_2, \ldots\}$ is the set of possible values of X.

Given a CDF $F_X(x)$, the PMF is determined by:

$$P_X(k) = F_X(k) - \lim_{x
ightarrow k^-} F_X(x) \quad = ext{ "jump" at } k$$

06 Illustration

\equiv Example - PDF and CDF: Roll 2 dice

Roll two dice colored red and green. Let X_R record the number of dots showing on the red die, X_G the number on the green die, and let S be a random variable giving the total number of dots showing after the roll, namely $S = X_R + X_G$.

- Find the PMFs of X_R and of X_G and of S.
- Find the CDF of *S*.
- Find P[S=8].

Solution

- 1. \equiv Sample space.
 - Denote outcomes with ordered pairs of numbers (*i*, *j*), where *i* is the number showing on the red die and *j* is the number on the green one.
 - Require that $i, j \in \mathbb{N}$ are integers satisfying $1 \leq i, j \leq 6$.
 - Events are sets of distinct such pairs.

2. \Rightarrow Create chart of outcomes.

Chart:									
+	1	2	3	4	5	6			
1	2	3	4	5	6	7			
2	3	4	5	6	7	8			
3	4	5	6	7	8	9			
4	5	6	7	8	9	10			
5	6	7	8	9	10	11			
6	7	8	9	10	11	12			

3. \equiv Definitions of X_R , X_G , and S.

• We have
$$X_R(i,j) = i$$
 and $X_G(i,j) = j$.

• Therefore S(i, j) = i + j.

4. \Rightarrow Find PMF of X_R .

• Use variable *n* for each possible value of
$$X_R$$
, so $n = 1, 2, ..., 6$

• Find $P_{X_R}(n)$:

$$P_{X_R}(n) \quad \gg \gg \quad P[X_R=n]$$

$$\gg \gg \frac{|\text{outcomes with } n \text{ on red}|}{|\text{all outcomes}|} \gg \gg \frac{6}{36} = \frac{1}{6}$$

• Therefore $P_{X_R}(n) = 1/6$ for every n.

5. \equiv Find PMF of X_G .

• Same as for X_R :

$$P_{X_G}(n)=rac{1}{6} \quad ext{for all } n$$

6. \mathbf{E} Find PMF of S.

• Find $P_S(n)$:

$$P_S(n) \quad \gg \gg \quad P[S=n] \quad \gg \gg \quad rac{| ext{outcomes with sum } n}{| ext{all outcomes}|}$$

• \triangle Count outcomes along *diagonal lines* in the chart.

• Create table of $P_S(n)$:

k	2	3	4	5	6	7	8	9	10	11	12
$p_S(k) = P(S = k)$	$\frac{1}{36}$	$\frac{2}{36}$	$\frac{3}{36}$	$\frac{4}{36}$	$\frac{5}{36}$	$\frac{6}{36}$	$\frac{5}{36}$	$\frac{4}{36}$	$\frac{3}{36}$	$\frac{2}{36}$	$\frac{1}{36}$



 \equiv Example - Total heads count; binomial expansion of 1

A fair coin is flipped n times.

Let *X* be the random variable that counts the total number of heads in each sequence.

The PMF of X is given by:

$$P_X(k) = {n \choose k} igg(rac{1}{2}igg)^r$$

Since the total probability must add to 1, we know this formula must hold:

$$egin{aligned} 1 &= \sum_{ ext{possible } k} P_X(k) \ &\gg & 1 &= \sum_{k=0}^n inom{n}{k} inom{1}{2}^n \end{aligned}$$

Is this equation really true?

There is another way to view this equation: it is the binomial expansion $(x + y)^n$ where $x = \frac{1}{2}$ and $y = \frac{1}{2}$:

$$\left(\frac{1}{2} + \frac{1}{2}\right)^n = \sum_{k=0}^n \binom{n}{k} \left(\frac{1}{2}\right)^n$$

A life insurance company has two clients, A and B, each with a policy that pays \$100,000 upon death. Consider events D_1 that the older client dies next year, and D_2 that the younger dies next year. Suppose $P[D_1] = 0.10$ and $P[D_2] = 0.05$.

Define a random variable X measuring the total money paid out next year in units of \$1,000. The possible values for X are 0, 100, 200. We calculate:

$$\begin{split} P[X=0] & \gg \gg \quad P[D_1^c]P[D_2^c] = 0.95 \cdot 0.90 = 0.86 \\ P[X=100] & \gg \gg \quad 0.05 \cdot 0.90 + 0.95 \cdot 0.10 = 0.14 \\ P[X=200] & \gg \gg \quad 0.05 \cdot 0.10 = 0.005 \end{split}$$